

A COUPLING FOR PRIME FACTORS OF A RANDOM INTEGER

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Joint work with Dimitris Koukoulopoulos

LARGE PRIME FACTORS

- Pick an integer N uniformly in $[1, x]$.
- Let $P_1 P_2 \cdots$ be the unique factorization of N with $P_1 \geq P_2 \geq \cdots$ being all primes or ones.

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THEOREM (BILLINGLSEY, 1972)

For any fixed $r \geq 1$, the random vector

$$\left(\frac{\log P_1}{\log x}, \dots, \frac{\log P_r}{\log x} \right) \xrightarrow{d} (V_1, \dots, V_r)$$

where $\mathbf{V} := (V_1, V_2, \dots)$ satisfy a **Poisson-Dirichlet distribution**.

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1. With the density of the first r components:

$$f(t_1, \dots, t_r) = \frac{1}{t_1 \cdots t_r} \cdot \rho \left(\frac{1 - (t_1 + \cdots + t_r)}{t_r} \right)$$

for $t_1 \geq t_2 \geq \cdots \geq t_r \geq 0$ and $\sum_i t_i \leq 1$.

The function $\rho: \mathbb{R}_{\geq 0} \rightarrow \mathbb{R}$ is the **Dickman-de Bruijn function**, characterized by:

- ▶ ρ is continuous;
- ▶ $\rho(u) = 1$ if $u \in [0, 1]$;
- ▶ $u\rho'(u) = -\rho(u-1)$ for $u \geq 1$.

2. With a stick-breaking process:

POISSON-DIRICHLET DISTRIBUTION

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- ▶ Sample $(U_i)_{i \geq 1}$ independently and uniformly in $[0, 1]$.

- ▶ Define

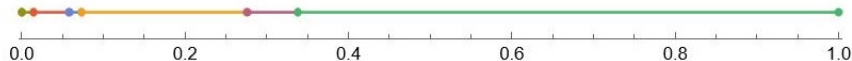
- ★ $L_1 := U_1$;

- ★ $L_i := U_i(1 - U_{i-1}) \cdots (1 - U_1)$ for $i \geq 2$.

GEM distribution : Distribution of $\mathbf{L} = (L_1, L_2, \dots)$.

- ▶ Let $\mathbf{V} = (V_1, V_2, \dots)$ be the non-increasing order statistics of \mathbf{L} .

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THEOREM (ARRATIA, 2002)

$$\mathbb{E} \sum_{i \geq 1} |\log P_i - (\log x) V_i| \ll \log \log x.$$

EXPECTED ℓ^1 -DISTANCE

DEFINITION

Let X and Y be two random variables on different probability spaces.

A **coupling** is a probability space over which there exists X' and Y' which have the same distribution as X and Y respectively.

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There exists a coupling of N and \mathbf{V} satisfying

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He conjectured that there exists a coupling such that this expectation is $O(1)$.

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We cannot do better than $O(1)$ because $\mathbb{E}[\#\{i \geq 1 : V_i \in [a, 2a]\}] = \log 2$ for all $a \in (0, \frac{1}{2})$.

Choose $a = \frac{\log 2}{3 \log x}$, then $(\log x) V_i \in [a, 2a]$ are all $\frac{\log 2}{3}$ away from any $\log p$.

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Possible to choose Q_i 's such that

$$\mathbb{P}[M \in A] = \frac{|A|}{x} + O\left(\frac{\log \log x}{\log x}\right) \text{ for any } A \subseteq \mathbb{N} \cap [1, x].$$

Arratia's original integer:

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MODIFYING ARRATIA'S COUPLING

Our modification:

- Choose $Q_i \in \{\text{prime powers}\} \cup \{1\}$ such that it is close to x^{L_i} .
- Set $J := \prod_{i \geq 2} Q_i$.
- Pick P_{extra} randomly with $\mathbb{P}[P_{\text{extra}} = p | J = j] = \frac{\log p}{\theta(x/j)}$.
- Set $M := J \cdot P_{\text{extra}}$.

Possible to choose Q_i 's such that

$$\mathbb{P}[M \in A] = \frac{|A|}{[x]} + O\left(\frac{1}{\log x}\right).$$

CONSTRUCTING M



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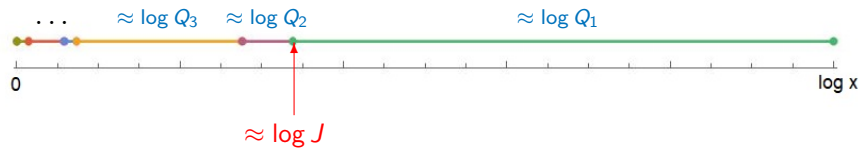
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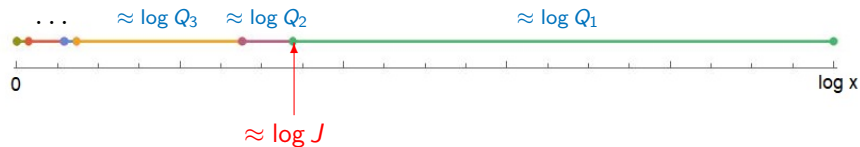
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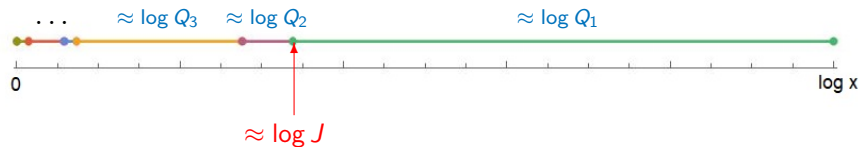


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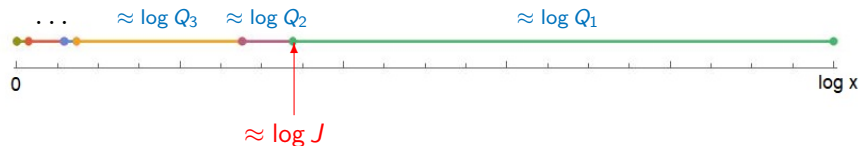
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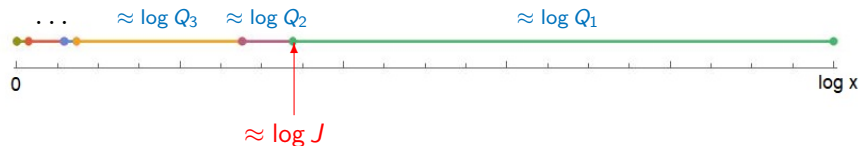
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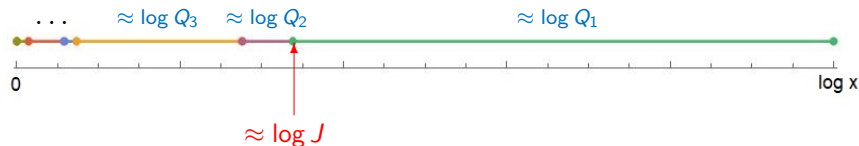
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- For fixed d , we have $\mathbb{P}[d|J] \rightarrow \frac{1}{d}$ as $x \rightarrow \infty$. ✓
- But J is **far** from uniform ($\mathbb{P}[J \leq y] \approx \frac{\log y}{\log x}$) ✗
- So we need to insert a minimal number of extra primes to **adjust** the distribution. This is why we need an extra prime!

THE TOTAL VARIATION DISTANCE

The **total variation distance** between the laws of two real-valued random variables X and Y is

$$d_{\text{TV}}(X, Y) := \sup_{A \subseteq \mathbb{R}} |\mathbb{P}[X \in A] - \mathbb{P}[Y \in A]|.$$

An equivalent definition is

$$d_{\text{TV}}(X, Y) := \min \mathbb{P}[X \neq Y]$$

where the minimum is over all possible couplings of X and Y .

Recall, that we have $\mathbb{P}[M \in A] = \mathbb{P}[N \in A] + O(\frac{1}{\log x})$,

This implies we can construct a random integer M from N so that $\mathbb{P}[M \neq N] \ll \frac{1}{\log x}$.

EXPECTED ℓ^1 -DISTANCE

Let \mathcal{E} be the event $M = N$ inside the coupling.

$$\mathbb{E} \sum_{i \geq 1} |\log P_i - (\log x) V_i| \leq \mathbb{E} \left[\mathbf{1}_{\mathcal{E}} \cdot \sum_{i \geq 1} |\log P_i - (\log x) V_i| \right] + (\log x) \cdot \mathbb{P}[\mathcal{E}^c]$$

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The P_i 's are now the prime factors of M inside \mathcal{E} . We reduce this problem to bounding

$$\mathbb{E} \sum_{i \geq 2} |\log Q_i - (\log x) L_i| + \mathbb{E}[\log P_{\text{extra}} - (\log x) L_1] + 2 \cdot (\log x) \cdot \mathbb{P}[\mathcal{E}^c].$$

We can compute that this is $O(1)$.

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PREDICTING THE SIZE OF DIVISORS

Here are the **divisors** on a **logarithmic scale** of two distinct integers around 3.52155×10^{25} with exactly 32 divisors:

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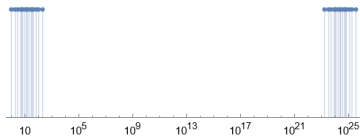


FIGURE: Divisors of $2 \times 3 \times 5 \times 7 \times 167693089728691213172671$

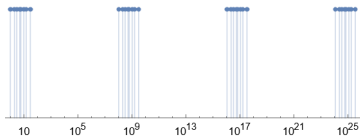


FIGURE: Divisors of $2 \times 3 \times 5 \times 105488303 \times 11127787008134317$

AVERAGE DISTRIBUTION OF DIVISORS

- Let N be a random integer uniform in $\mathbb{N} \cap [1, x]$.
- Let D be a random divisor of N uniformly distributed.

THEOREM

For $x \geq 2$ and uniformly for $u \in [0, 1]$, we have

$$\mathbb{P}[D < N^u] = \frac{2}{\pi} \arcsin(\sqrt{u}) + \dots$$

- *Deshouillers, Dress, Tenenbaum (1979):*
 $\dots + O\left(\frac{1}{\sqrt{\log x}}\right)$ for $u \in [0, 1]$.
- *Bareikis, Manstavičius (2007):*
 $\dots + O\left(\frac{1}{\sqrt{u(1-u)} \log x}\right)$ for $u \in \left[\frac{1}{\log x}, 1 - \frac{1}{\log x}\right]$.

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Estimated $\mathbb{P}[D < N^u]$ when $\mathbb{P}[D = d \mid N = n] = f(d)$ for any $d|n$, and f being **multiplicative, fixed at primes**, satisfying $1 * f = 1$ and $f \geq 0$.

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Estimated $\mathbb{P}[D_1 < N^{u_1} \text{ and } D_2 < N^{u_2}]$ when D_1 is a random divisor (unif. dist.), and D_2 is a random divisor of $\frac{N}{D_1}$ (unif. dist.).

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- *Leung (2023)*:

Estimated $\mathbb{P}[D_1 < N^{u_1}, \dots, D_{k-1} < N^{u_{k-1}}]$ for any k , for $\mathbb{P}[D_1 = d_1, \dots, D_k = d_k \mid N = n] = f(d_1, \dots, d_k)$ if $d_1 \cdots d_k = n$, and f is **multiplicative, fixed at primes**.

The proofs of these theorems all involve getting asymptotics for sums of multiplicative functions (Landau–Selberg–Delange method).

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Q: Can we use the distribution of large prime factors to get a similar result?

AVERAGE DISTRIBUTION OF DIVISORS

Probabilistic version:

THEOREM (DONNELLY AND TAVARÉ, 1987)

- χ random k -colouring of \mathbb{N} , with $\mathbb{P}[\chi(i) = j] = \frac{1}{k}$.
- $Z_m := \sum_{i \geq 1} 1_{\chi(i)=j} V_i$ for $1 \leq m \leq k$.

Then $\mathbf{Z} := (Z_1, \dots, Z_k) \sim \text{Dir}(\frac{1}{k}, \dots, \frac{1}{k})$.

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Arratia (1998): Used probabilistic methods to prove $\mathbb{P}[D < N^u] = \frac{2}{\pi} \arcsin \sqrt{u} + o(1)$.

There exists a prob. space where $V_i \approx \frac{\log P_i}{\log x}$, we assign a random colour to each prime factor. This creates $\mathbf{D} := (\frac{\log D_1}{\log N}, \dots, \frac{\log D_k}{\log N})$. We have $\|\mathbf{D} - \mathbf{Z}\|_\infty \asymp \frac{1}{\log x}$ typically.

FACTORIZATIONS INTO k PARTS

THEOREM (LEUNG, 2023)

Let $\mathbf{u} \in [0, 1]^{k-1}$ with $u_1 + \dots + u_{k-1} \leq 1$.

Pick N uniformly in $[1, x]$. Pick $D_1 \cdots D_k = N$ uniformly among all k -factorizations, then

$$\begin{aligned} \mathbb{P}[D_1 < N^{u_1}, \dots, D_{k-1} < N^{u_{k-1}}] \\ = F(u_1, \dots, u_{k-1}) + O_k \left(\frac{1}{(\log x)^{1/k}} \right). \end{aligned}$$

Used variant of Landau-Selberg-Delange method

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Error term: Deviation by $\asymp \frac{1}{\log x}$ from u_i of a $\text{Dir}(\frac{1}{k}, \dots, \frac{1}{k})$.

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Pick N uniformly in $[1, x]$. Pick $D_1 \cdots D_k = N$ uniformly among all k -factorizations, then

$$\begin{aligned} \mathbb{P}[D_1 < N^{u_1}, \dots, D_{k-1} < N^{u_{k-1}}] \\ = F(u_1, \dots, u_{k-1}) + O\left(\sum_{1 \leq i < k} \frac{1}{\xi(u_i, x) \log x}\right) \end{aligned}$$

with $\xi(u, x) = \max\{u, \frac{1}{\log x}\}^{1-1/k} \cdot \max\{1-u, \frac{1}{\log x}\}^{1/k}$.

Error term: Deviation by $\asymp \frac{1}{\log x}$ from u_i of a $\text{Dir}(\frac{1}{k}, \dots, \frac{1}{k})$.

FACTORIZATIONS INTO k PARTS

THEOREM (H., KOUKOULOPOULOS, 2024)

Let $\mathbf{u} \in [0, 1]^{k-1}$.

Pick N uniformly in $[1, x]$. Pick $D_1 \cdots D_k = N$ uniformly among all k -factorizations, then

$$\begin{aligned} \mathbb{P}[D_1 < N^{u_1}, \dots, D_{k-1} < N^{u_{k-1}}] \\ = F(u_1, \dots, u_{k-1}) + O\left(\sum_{1 \leq i < k} \frac{1}{\xi(u_i, x) \log x}\right) \end{aligned}$$

with $\xi(u, x) = \max\{u, \frac{1}{\log x}\}^{1-1/k} \cdot \max\{1-u, \frac{1}{\log x}\}^{1/k}$.

Error term: Deviation by $\asymp \frac{1}{\log x}$ from u_i of a $\text{Dir}(\frac{1}{k}, \dots, \frac{1}{k})$...

... and the fact that $(\frac{\log D_1}{\log N}, \dots, \frac{\log D_k}{\log N})$ has discrete support (e.g.: $u_i < \frac{\log 2}{\log x}$ and $D_i < N^{u_i}$ implies $D_i = 1$).



- Sizes of the prime factors of N converge to components of \mathbf{V}



- Sizes of the prime factors of N converge to components of \mathbf{V}
- Construct a space with both N and \mathbf{V} , such that prime factors of N are close to components of \mathbf{V} most of the time
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- Construct a space with both N and \mathbf{V} , such that prime factors of N are close to components of \mathbf{V} most of the time
- Use this space as a bridge to pass from probabilistic model to get to information on stats of divisors!